Remote Vector Gaussian Source Coding With Decoder Side Information Under Mutual Information and Distortion Constraints

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Abstract—Let $X$, $Y$, $Z$ be zero-mean, jointly Gaussian random vectors of dimensions $n_x$, $n_y$, and $n_z$, respectively. Let $\mathcal{P}$ be the set of random variables $W$ such that $W \leftrightarrow Y \leftrightarrow (X, Z)$ is a Markov string. We consider the following optimization problem:

$$\min_{W \in \mathcal{P}} I(Y; W | Z)$$

subject to one of the following two possible constraints: 1) $I(X; W | Z) \geq R_x$, and 2) the mean squared error between $X$ and $\hat{X} = \mathbb{E}(X | W, Z)$ is less than $d$. The problem under the first kind of constraint is motivated by multiple-input multiple-output (MIMO) relay channels with an oblivious transmitter and a relay connected to the receiver through a dedicated link, while for the second case, it is motivated by source coding with decoder side information where the sensor observation is noisy. In both cases, we show that jointly Gaussian solutions are optimal. Moreover, explicit water filling interpretations are given for both cases, which suggest transform coding approaches performed in different transform domains, and that the optimal solution for one problem is, in general, suboptimal for the other.

Index Terms—Gaussian vector, relay channel, source coding with side information, transform coding.

I. INTRODUCTION

In a relay channel [1], the relay wishes to help the transmitter to convey information reliably to the end receiver. In certain cases, the transmitter does not actively cooperate with the relay, which falls into the oblivious cooperation framework [2]; it can also be understood as a special case of the nomadic transmitter setting where two relays send summary information regarding their received signals to a final central processing center [3], but in this special case, the link between one relay and the processing center is taken to have an almost infinite capacity. In both settings, the transmitter may send its signal without awareness of the existence of the relay, and the relay does not necessarily know exactly the codebooks the transmitter is using.

We are interested in the case that the relay is connected to the receiver through a dedicated link (e.g., a separate frequency band or a land line) of a maximum rate $R$. We assume the channel noises are additive and Gaussian, and moreover, the transmitter which has $n_t$ transmit antennas is using Gaussian signaling $X$; this is a reasonable and popular assumption [3]. The relay which has $n_r$ receive antennas observes $Y$, and it is a noisy version of $X$. The relay now has the task of source coding this noisy observation for the end receiver. The end receiver, which is the decoder in this source coding setting, has $n_z$ antennas and also has side information $Z$ through its own observation of the channel. Assuming channel state information and the probability distribution of $X$ are known at the relay and end receiver, and using well-known properties of the typical sequences, it can be shown that the system can reliably transmit at rate arbitrarily close to $R_{T^d}$, if there exists a random variable $W \leftrightarrow Y \leftrightarrow (X, Z)$, such that

$$I(X; Z, W) \geq R_{T^d} \quad I(Y; W | Z) \leq R$$

This can be rewritten as the optimization problem

$$R(R_t) \triangleq \min_{W \in \mathcal{P}} I(Y; W | Z)$$

subject to

$$I(X; W | Z) = I(X; Z, W) - I(X; Z) \geq R_{T^d} - I(X; Z) \triangleq R_t$$

which is exactly the optimization problem under the first kind of constraint given in the abstract. Roughly speaking, the relay utilizes a source coding scheme based on the Wyner–Ziv coding technique [4], and the quantity $I(Y; W | Z)$ is exactly this coding rate, however the receiver places the mutual information constraint (3), instead of attempting to reconstruct the source $X$ to satisfy certain distortion constraint as in [4].

Let us turn to the more conventional rate-distortion setting where the encoder observes a noisy version of the source, and the end receiver, which has certain side information, is interested in recovering the source within certain distortion. Wyner and Ziv’s solution [4] generalizes naturally to this remote source setting (see [5] and [6]). The single-letter optimization posed in [1].

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the abstract under the second constraint is thus exactly the rate distortion function under a mean squared error (MSE) distortion criterion, i.e., the quantity $I(Y; W | Z)$ is again the source coding rate. We will denote this function as $R(d)$. It is now clear that the essential task is to determine the optimal random variable $W$, since in both problems it provides the optimal quantization codebook, and perhaps more importantly, the optimal resource allocation for these vector source coding problems. One key question in this context is whether the optimal solutions of the random variables $W$ are the same under these two constraints.

In this paper, we show that for both cases a jointly Gaussian vector solution is optimal for $W$, which also naturally leads to transform coding and water-fill interpretations. However, the water filling is performed in different spaces for the two problems, which implies that, in general, the optimal solutions in the rate distortion setting are suboptimal solutions in the relay setting, and *vice versa*. It is thus important to distinguish between these two constraints, and use the appropriate one depending on the particular system requirement.

One may also ask whether giving the decoder side information $Z$ to the encoder helps. Indeed our result shows that encoder side information does not help for both cases, similar to the well-known result in Wyner–Ziv source coding of scalar Gaussian sources.

The solution of the vector problem under the mutual information constraint is largely inspired by the work in [8] for the case without side information at the decoder. However, our proof is significantly simpler than theirs, despite the fact that side information is introduced. Our work is also relevant to [7]. However, in [7], only distortion-constrained problem with direct observation of the source is considered, which does not include the problems considered here.

In the rest of this paper, the following notation is adopted. Random vectors are written in bold capital letters, and random variables in capital letters. $X, Y, Z$ are jointly Gaussian, zero-mean, real-valued random vectors. We write their variance matrices as $\Sigma_{xx}, \Sigma_{yy}, \Sigma_{zz}$, and covariance matrices as $\Sigma_{xy}, \Sigma_{yz}, \Sigma_{zx}$, respectively. $X$ is the underlying source of interest, $Y$ is the noisy observation at the encoder, and $Z$ is the decoder side information. We do not assume that $Y$ and $Z$ are conditionally independent given $X$, i.e., the channel noises are not necessarily independent. All the logarithms are base $e$.

Before proceeding, we will rewrite the relation between $X$, $Y$, and $Z$ to simplify the proof given in the sequel. It is clear that without loss of generality, we can write

$$\begin{align*}
Y &= K_{y|x}Z + N_y \\
X &= K_{x|y}Y + K_{x|z}Z + N_x \\
&= (K_{x|y}K_{y|x} + K_{x|z})Z + K_{x|y}N_y + N_x
\end{align*}$$

(4)

where $N_x$ and $N_y$ are zero-mean jointly Gaussian random vectors independent of each other, $N_x$ is independent of $Z$, and $N_y$ is independent of $(Y, Z)$. More precisely, we have by applying well-known linear estimation calculation [9] that

$$K_{y|x} = \Sigma_{y|x}^{-1}\Sigma_{y|x}$$

$$K_{x|y} = \Sigma_{x|y}^{-1} + \Sigma_{x|y}^{-1}\Sigma_{y|x}\Delta^{-1}\Sigma_{y|x}^{-1} - \Sigma_{x|y}^{-1}\Sigma_{y|x}\Delta^{-1}$$

(5)

where $\Delta = \Sigma_{x} - \Sigma_{y|x}\Sigma_{y|x}^{-1}\Sigma_{y|x}$. We will assume the inverses of the above matrices are well defined, which is usually satisfied except for some degenerate cases. The following covariance matrices are also important:

$$\begin{align*}
\Sigma_{N_1} &= \Sigma_{y|x} - K_{y|x}\Sigma_{y|x} \\
\Sigma_{N_2} &= \Sigma_{x|y} - K_{x|y}\Sigma_{x|y} \\
\Sigma_{x|z} &= \Sigma_{x|z} - \Sigma_{x|y}\Sigma_{y|x}\Delta^{-1}\Sigma_{y|x}
\end{align*}$$

(6)

II. SCALAR RANDOM VARIABLE PROBLEMS

For the scalar case, denote $k_{y|x} \triangleq K_{y|x}, k_{x|y} \triangleq K_{x|y},$ and $k_{x|z} \triangleq K_{x|z}$, respectively. We consider the scalar Gaussian case first. Since we are not aware of a formal proof for the case under the mutual information constraint, such a proof is given; for the other case, we simply restate the known result.

A. The Mutual Information Constraint Problem

*Theorem 1*: For the scalar Gaussian source and $R_I \geq 0$

$$R(R_I) = \infty \quad \text{if} \quad R_I \geq \frac{1}{2} \log \frac{\Sigma_{x|z}}{\Sigma_{x|y}}$$

(7)

and, otherwise

$$R(R_I) = \frac{1}{2} \log \frac{\Sigma_{x|z} - \Sigma_{x|y}}{\Sigma_{x|z} \exp(-2R_I) - \Sigma_{x|y}}.$$  

(8)

*Proof*: We start the converse proof by writing

$$I(X; W | Z) = h(X|Z) - h(X|W, Z).$$

(9)

To bound the second term, we have

$$h(X|W, Z) = h(k_{x|y}Y + k_{x|z}Z + N_x|W, Z) = h(k_{x|y}Y + N_x|W, Z).$$

(10)

Then, it follows:

$$\exp[2h(k_{x|y}Y + N_x|W, Z)]$$

$$\geq \exp[2h(k_{x|y}Y|W, Z)] + \exp[2h(N_x)]$$

$$= k_{x|y}^2 \exp[2h(Y|W, Z)] + \exp[2h(N_x)]$$

$$= k_{x|y}^2 \exp[-2I(Y; W|Z) + 2h(Y|Z)] + \exp[2h(N_x)]$$

$$= k_{x|y}^2 \exp[-2I(Y; W|Z) + 2h(N_x)] + \exp[2h(N_x)]$$

(11)

where (a) is by applying the conditional version of the entropy power inequality [10]. Thus, we have

$$I(Y; W | Z) \geq \frac{1}{2} \log[k_{y|x}^2 \exp[2h(N_1)]]$$

$$\geq \frac{1}{2} \log[\exp[2h(X|Z)] - 2I(Y; W|Z) - \exp[2h(N_2)]].$$

(12)
Because the function \( -\log(r \exp(-2x) - q) \) is monotonically increasing in \( x \) when \( r > 0 \), and \( I(W; X|Z) \geq R_f \), we have

\[
I(Y; W|Z) \geq \frac{1}{2} \log \left[ \frac{\sum_{x|y} \exp(2h(N_1))}{\exp(2h(N_2))} \right] - \frac{1}{2} \log \left[ \frac{\exp(2h(X|Z) - 2R_f)}{\exp(2h(N_2))} \right].
\]

(12)

Clearly, when we take \( W = Y + N_3 \) where \( N_3 \) is a zero-mean Gaussian independent of everything else such that \( I(W; X|Z) = R_f \), all the inequalities become equality.

Since \( I(W; X|Z) \leq I(W; Y; X|Z) \), the Markov string \( W \leftrightarrow Y \leftrightarrow (X, Z) \) implies that \( I(W; X|Z) = I(Y; X|Z) \). Thus, when \( R_f \geq I(Y; X|Z) \), the optimization problem is not feasible. Substitute appropriate values into the quantities, and it is easy to see \( \sum_{x|y} - \sum_{x|y} = k_{xy}^2 \sum_{y|z} \) because it is the variance of \( k_{xy}N_1 \). The proof is thus complete. \( \square \)

B. The Rate-Distortion Problem

**Theorem 2** ([15], [6]): For the scalar Gaussian source, we have for \( \sum_{x|y} \geq d \geq \sum_{x|y} \)

\[
R(d) = \frac{1}{2} \log \left[ \frac{\sum_{x|y} - \sum_{x|y}}{d - \sum_{x|y}} \right]
\]

and \( R(d) = 0 \) if \( d > \sum_{x|y} \); \( R(d) = \infty \) if \( 0 \leq d < \sum_{x|y} \).

We omit the proof here; see [5] and [6]. It can be shown that in both cases encoder side information does not help, i.e., even when the Markov string is relaxed to \( W \leftrightarrow (Z, Y) \leftrightarrow X \), the solutions do not change. Further, notice the similarity between the functions \( R(R_f) \) and \( R(d) \). Indeed, for the scalar problem, for both cases, the solution is given by the same Gaussian solution \( W = Y + N_3 \) with the variance of \( N_3 \) properly chosen. However, as we will illustrate next, for the vector problem, their solutions no longer coincide.

III. RANDOM VECTOR PROBLEMS

In this section, we discuss the vector version of the problems.

We provide a proof by decomposing the problem into a set of componentwise problems. It is conceivable that the enhancement technique given in [11] and [12] may be of use, however our proof directly yields the water filling interpretation.

A. The Mutual Information Constraint Problem

The following notation is needed. For a symmetric positive-definite matrix \( S \), let its eigenvalue decomposition be \( U^T \Lambda U \), where \( \Lambda \) is the diagonal matrix with eigenvalues on its diagonal. Define \( S^{1/2} = U^T \Lambda^{1/2} U \), where \( \Lambda^{1/2} \) is the diagonal matrix where the square root of each element of \( \Lambda \) is taken. Similarly, \( S^{-1/2} \) is also defined for a symmetric positive-definite matrix, however with an inverse on each nonzero element of \( \Lambda \). The main result in this section is the following theorem.

**Theorem 3**: Let the eigenvalue of \( \sum_{x|y} k_{xy}^T \sum_{y|z} k_{yx} \sum_{y|z} \) be given as \( \mu_i \). Then, for \( 0 \leq R_f \leq \frac{1}{2} \log \left[ \frac{\sum_{x|y} - \sum_{x|y}}{\mu_i} \right] \),

\[
R(R_f) = \sum_{i=1}^{n_z} \frac{1}{2} \log \left( \frac{\mu_i}{1 - \gamma} \right)^{-1} - (1 - \mu_i)^{-1}
\]

where \( (x)^{-} = \min(x, 1) \), \( \log_0 0 = 0 \), and \( \gamma \in [0, 1) \) is chosen such that

\[
-\frac{1}{2} \sum_{i=1}^{n_z} \log \left( \frac{1 - \mu_i}{1 - \gamma} \right) = R_f.
\]

If \( R_f > R_f^* \), then \( R(R_f) = \infty \).

We will consider a specific transform; this transform is usually referred to as the canonical correlation analysis in statistics. First, find the singular value decomposition of \( \sum_{x|y} = \sum_{y|z} = U^T \Lambda U \), where \( U \) and \( \Lambda \) are orthogonal matrices and \( \Lambda \) is a diagonal matrix with singular values on the diagonal. Define \( T_x = U \sum_{x|y}^{-1/2} \) and \( T_y = V \sum_{y|z}^{-1/2} \), which in fact gives the transforms \( X' = T_x X \) and \( Y' = T_y Y \), respectively. Notice that the transformation matrices are full rank, and under these transforms

\[
\sum_{x|y} = T_x \sum_{x|y} T_x^T = U \sum_{x|y}^{-1/2} \sum_{x|y} \sum_{x|y}^{-1/2} U^T = I_{n_x}
\]

where \( I_{n_x} \) is the identity matrix of dimension \( n_x \). Similarly

\[
\sum_{y|z} = T_y \sum_{y|z} T_y^T = V \sum_{y|z}^{-1/2} \sum_{y|z} \sum_{y|z}^{-1/2} V^T = I_{n_y}
\]

Moreover, we have

\[
\sum_{x|y} \sum_{y|z} = T_x T_y \sum_{x|y} \sum_{y|z} T_y T_x^T = U \sum_{x|y}^{-1/2} \sum_{x|y} \sum_{y|z} \sum_{y|z}^{-1/2} V^T = \Lambda\Lambda \Gamma.
\]

These facts imply that after the transform, the components of \( X' \) and \( Y' \), respectively, are independent and identically distributed with unit variance given \( Z \). The vectors of \( X' \) and \( Y' \) are only componentwise dependent given \( Z \). Moreover, since the element of \( \sum_{x|y} \) is \( \Lambda \) is a covariance of two random variables of unit variance, its diagonal values satisfy \( 0 \leq |\lambda_i| \leq 1 \).

**Proof**: Since the transform \( T_x \) and \( T_y \) are full rank, and thus information lossless, we have

\[
I(X; W|Z) = I(X'; W|Z).
\]

Thus, we have by the conditional independence and the chain rule

\[
I(X; W|Z) = \sum_{i=1}^{n_z} [h(X'_i|Z) - h(X'_i|W, Z, X'_z)]
\]

\[
\leq \sum_{i=1}^{n_z} [h(X'_i|Z) - h(X'_i|W, Z, X'_r, Y'_r)]
\]

\[
\leq h(X'_i|Z) - h(X'_i|W, Z, Y'_r|Z)
\]

\[
= \sum_{i=1}^{n_z} I(X'_i; W; Y'_r|Z)
\]

(18)

where \( (a) \) is due to the Markov string \( X'_i \leftrightarrow (Y'_r, W, Z) \leftrightarrow X'_r \). On the other hand, we have

\[
R \geq I(Y; W|Z) = \sum_{i=1}^{n_y} [h(Y'_i|Z) - h(Y'_i|W, Z, Y'_r|Z)]
\]

\[
= \sum_{i=1}^{n_y} I(Y'_i; W, Y'_r|Z).
\]

(19)
Now define \((W, Y_{i}^{T}) \leftrightarrow W_{i}\), and it is clear that \(X_{i}^{T} \leftrightarrow (Z_{i}, Y_{i}^{T}) \leftrightarrow W_{i}\) is a Markov string. Thus, we have relaxed the problem into
\[
\min_{\lambda} \sum_{i=1}^{n_{y}} I(Y_{i}^{T}; W_{i}|Z)
\]
subject to: \(\sum_{i=1}^{n_{y}} I(X_{i}^{T}; W_{i}|Z) \geq R_{I}\) (21)

where \(X_{i}^{T} \leftrightarrow (Z_{i}, Y_{i}^{T}) \leftrightarrow W_{i}\) for \(i \leq \max(n_{x}, n_{y})\).

Using the same outer bounding derivation as the scalar problem, we can further write the relaxed problem as2
\[
\min_{\lambda} \sum_{i=1}^{n_{y}} -\frac{1}{2} \log[\exp(-2R_{i}) - (1 - \lambda_{i}^{2})] + \frac{1}{2} \log[\lambda_{i}^{2}]
\]
subject to: \(\sum_{i=1}^{n_{y}} R_{i} \geq R_{I}, -\frac{1}{2} \log(1 - \lambda_{i}^{2}) \geq R_{i} \geq 0\) (23)

where \(n_{I}\) is the number of nonzero \(\lambda_{i}\) values.

A water-filling solution can be derived from the above optimization problem. More precisely, by the Kuhn–Tucker condition, we can find the optimizing solution as
\[
R_{i} = \frac{1}{2} \log \left( \frac{\lambda_{i} - 1}{\lambda_{i}(1 - \lambda_{i}^{2})} \right)^{+}
\]
where we use \((x)^{+}\) to denote \(\max(x, 1)\), and \(\lambda_{i}\) is the “water level” such that the sum-rate constraint is satisfied with equality. Now defining \(\gamma = \lambda^{-1}\) and noticing \(\lambda_{i}^{2} = \mu_{i}\) give the formula in the theorem.

It remains to show that the solution of the relaxed problem is indeed achievable, i.e., there exists a random variable (vector) \(W\) that matches this lower bound, but this is obvious by letting \(W = T_{1} Y + N_{3}\), where \(N_{3}\) is a Gaussian random vector with independent components such that the equality \(I(X_{i}^{T}; Y_{i}^{T} + N_{3}; W) = R_{i}\) holds. This indeed is a Gaussian solution, and the theorem is proved. \(\Box\)

B. The Rate-Distortion Problem

Assuming the MSE distortion is taken to be averaged over all components, we have the following theorem.

**Theorem 4:** Let the eigenvalues of \(K_{xy}^{T} N_{y|x} K_{xy}\) be given as \(\lambda_{i}^{D}, i = 1, 2, \ldots, n_{x}\). Then, for \(\text{Tr}(\Sigma_{y|x}) \geq n_{x} d \geq \text{Tr}(\Sigma_{y|x|yz})\)
\[
R(d) = \sum_{i=1}^{n_{x}} \frac{1}{2} \log \left( \frac{\lambda_{i}^{D}}{\lambda} \right)^{+}
\]
(24)

where \(\text{Tr}(\cdot)\) is the trace of a matrix, and \(\lambda > 0\) is chosen such that
\[
\sum_{i=1}^{n_{x}} \min(\lambda, \lambda_{i}^{D}) = n_{x} d - \text{Tr}[\Sigma_{y|x|yz}].
\]
(25)

If \(n_{x} d < \text{Tr}(\Sigma_{y|x|yz})\), then \(R(d) = \infty\); if \(n_{x} d > \text{Tr}(\Sigma_{y|x|yz})\), then \(R(d) = 0\).

The proof of this theorem is rather standard, however we include it below for completeness.

**Proof:** To prove the converse, consider the fact that
\[
\text{Tr}[E(\hat{X} - \hat{X})(\hat{X} - \hat{X})^{T}] \leq n_{x} d,
\]
(26)

It is clear that
\[
E(\hat{X} - \hat{X})(\hat{X} - \hat{X})^{T} \nonumber
= E(K_{xy} Y + K_{xz} Z - \hat{X})(K_{xy} Y + K_{xz} Z - \hat{X})^{T} + \Sigma_{y|x|yz}
\]
(27)

because \(N_{2}\) is independent of \(Y\) and \(Z\), thus independent of \(W\) and \(Z\). Now define \(\tilde{X} = \hat{X} + (K_{xy} Y + K_{xz} Z)\), from which we see that
\[
\text{Tr}\left[ E(\hat{X} - \hat{X})(\hat{X} - \hat{X})^{T} \right] 
= \text{Tr}\left[ E(K_{xy} N_{y} - \tilde{X})(K_{xy} N_{y} - \tilde{X})^{T} \right]
+ \text{Tr}[\Sigma_{y|x|yz}] \leq n_{x} d.
\]
(27)

Thus, there exists a function \(\hat{X} = f(W, Z)\) such that
\[
\text{Tr} \left[ E(K_{xy} N_{y} - \tilde{X})(K_{xy} N_{y} - \tilde{X})^{T} \right]
\leq n_{x} d - \text{Tr}[\Sigma_{y|x|yz}] \triangleq n_{x} d'.
\]
(28)

Denote the eigenvalue decomposition of \(K_{xy}^{T} N_{y|x} K_{xy}\) as \(U_{xy}^{T} \Lambda U_{xy}\). Denote \(N_{y}^{T} = U_{xy} K_{xy} N_{y}\), which clearly has independent components. We have the following chain of inequalities:
\[
I(Y; W|Z) \geq I(U_{xy} Y; W|Z) 
= h(U_{xy} Y|Z) - h(U_{xy} Y|W, Z)
= h(U_{xy} N_{y}) - h(U_{xy} N_{y}|W, Z)
\geq \sum_{i=1}^{n_{y}} h(N_{y}^{(i)}) - h(N_{y}^{(i)}|W, Z, N_{y}^{(i)})
\geq \sum_{i=1}^{n_{y}} I(N_{y}^{(i)}; W, Z)
\geq \sum_{i=1}^{n_{y}} I(N_{y}^{(i)}; E(N_{y}^{(i)}|W, Z)) = \sum_{i=1}^{n_{y}} I(N_{y}^{(i)}; \tilde{N}_{y}^{(i)})
\]
(29)

where \(\tilde{N}_{y}^{(i)} \triangleq E(N_{y}^{(i)}|W, Z)\). Notice that \(U\) is an orthogonal transform, thus preserves the MSE distortion. Define \(d_{i} = E(N_{y}^{(i)} - \tilde{N}_{y}^{(i)})^{2}\), then it is clear that we have \(\sum_{i=1}^{n_{y}} d_{i} \leq n_{x} d'\). This is a conventional rate-distortion problem, and thus
\[
R(d) \geq \sum_{i=1}^{n_{y}} \frac{1}{2} \log \left( \frac{\lambda_{i}^{D}}{\lambda} \right)^{+}
\]
(30)

where \(\lambda\) is the water level such that
\[
\sum_{i=1}^{n_{y}} \min(\lambda, \lambda_{i}^{D}) = n_{x} d - \text{Tr}[\Sigma_{y|x|yz}].
\]
(31)
This establishes the lower bound for the rate distortion function. The fact that this lower bound is achievable is straightforward by letting $W = U K_{xy} Y + N_3$, where $N_3$ is a Gaussian vector having independent components, with their variances determined by

$$\sigma_s^2 = \min(\lambda, \lambda_D) \frac{1}{\lambda - \min(\lambda, \lambda_D)},$$

(32)

The proof is complete. □

IV. CONCLUDING REMARKS

From the proofs of Theorem 3 and Theorem 4, it is seen that the two transforms in which water filling is performed are different under the two kinds of constraints. For the mutual information constraint, the transform is $T_y = V_2^{-1/2} K_{zy}^T$, where $V_2$ is determined by the eigenvalue decomposition of $\Sigma_{yy}^{-1/2} K_{xy}^T K_{xy} \Sigma_{yy}^{-1/2}$, and the water filling is along these eigenvalues. For the MSE distortion constraint, the transform is determined by the eigenvalue decomposition of $K_{xy} \Sigma_{yy}^{-1} K_{xy}^T$, and the water filling is along these eigenvalues. Moreover, the transform $T_y$ may not be orthogonal, thus it does not preserve trace and the MSE distortion. This result implies that the optimal solution for one problem is, in general, suboptimal for the other. Thus, when multiple-input multiple-output (MIMO) system is considered for the relay problem, it is important to recognize that the Wyner-Ziv source coding component should be used under the mutual information constraint, instead of distortion constraint. It should also be noted that the problem under mutual information constraint considered here includes the hypothesis testing problem in [13] as a special case when the source is vector Gaussian, and thus the given result provides an explicit solution for it.

It is theoretically interesting to consider the same optimization problem when the two kinds of constraints are placed simultaneously. The extremal inequality in [11] indicates that the optimal value is attained by timesharing Gaussian solutions. However, it is unclear whether a single Gaussian solution suffices and whether there exists an explicit water filling interpretation.

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